





HAFS as a Testbed for Non-Gaussian Data Assimilation Developments for the UFS

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Primary objective: Implement novel DA methodology that is immediately relevant for HAFS.

Specific topics:

New developments for fully cycled ensemble DA within HAFS

- * Bias correction for radiance measurements (Knisely and Poterjoy, 2023; UIFCW talk on Monday)
- * Treating sampling error in high-resolution ensembles (Kurosawa and Poterjoy, in progress)
- * Non-Gaussian errors (Poterjoy 2022a,b; Kurosawa and Poterjoy 2021,2023; UIFCW poster)

Broadly relevant to all UFS applications.

One objective is to explore implications of replacing the EnKF with LPF for modeling systems that run EnVar.

Motivation:

- Most modeling systems run EnVar for practical reasons; e.g., use of a high-resolution deterministic "control."
- EnKF is typically used to update ensemble—to provide future background error covariance for EnVar.
- EnKF members are re-centered on EnVar analysis.

One objective is to explore implications of replacing the EnKF with LPF for modeling systems that run EnVar.

Motivation:



- i. Posterior tends to be closer to a Gaussian than the prior.
- ii. Re-centering posterior ensemble on Var analysis is okay, as long as the distribution is close to Gaussian.

← Var analysis alongside PF members following assimilation.

One objective is to explore implications of replacing the EnKF with LPF for modeling systems that run EnVar.

Motivation:



- iii. Incremental 3DVar/4DVar can solve moderately nonlinear DA problems through an outer loop (e.g., x on left).
- iv. Posterior targeted by Var is more consistent with PF than EnKF.
- ← Var analysis alongside **EnKF members**.

Real-world impact of assuming Gaussian prior



Combining particle filters with Var

DA comparisons:

- "EnKF-Var" ← HAFS ensemble updated with EnKF and Var
- "PF-Var" ← HAFS ensemble updated with LPF and Var

In both experiments, role of EnKF or LPF is to update 40 HAFS ensemble members about a variational analysis.

Verification:

- 10-member forecasts generated every 6 h for 2 weeks
- Storm features verified using NHC Best Track data
- Synoptic scale features verified using ERA5

Verification (2 weeks of forecasts)



 Currently testing with 2023 HAFS-A and HAFS-B; preliminary results shows similar improvements with LPF-Var.

Verification (2 weeks of forecasts)



 LPF will soon be applied for hourly-updated GFS (FY23 WPO Innovations for Community Modeling Competition). Flexibility provided by non-Gaussian data assimilation:

 $p(\mathbf{x}_t|\mathbf{y}_{0:t}) \propto p(\mathbf{y}_t|\mathbf{x}_t)p(\mathbf{x}_t|\mathbf{y}_{0:t-1}),$

Flexibility provided by non-Gaussian data assimilation:

$$\begin{split} \rho(\mathbf{x}_t | \mathbf{y}_{0:t}) &\propto \quad \rho(\mathbf{y}_t | \mathbf{x}_t) \rho(\mathbf{x}_t | \mathbf{y}_{0:t-1}), \\ &\approx \quad \rho(\mathbf{y}_t | \mathbf{x}_t) \frac{1}{N_e} \sum_{n=1}^{N_e} \delta(\mathbf{x} - \mathbf{x}_t^n), \end{split}$$

Flexibility provided by non-Gaussian data assimilation:

$$p(\mathbf{x}_t | \mathbf{y}_{0:t}) \propto p(\mathbf{y}_t | \mathbf{x}_t) p(\mathbf{x}_t | \mathbf{y}_{0:t-1}),$$

$$\approx p(\mathbf{y}_t | \mathbf{x}_t) \frac{1}{N_e} \sum_{n=1}^{N_e} \delta(\mathbf{x} - \mathbf{x}_t^n),$$

$$\propto \sum_{n=1}^{N_e} p(\mathbf{y}_t | \mathbf{x}_t^n) \delta(\mathbf{x} - \mathbf{x}_t^n).$$

Large freedom exists in how we specify $p(\mathbf{y}_t | \mathbf{x}_t^n)$.

Assimilating obs with non-Gaussian, state-dependent errors.

- Model III of Lorenz (2005) on periodic domain
- Model configuration supports chaotic behavior
- Characterized by N_x = 480 variables on periodic domain
- Data Assimilation: <u>iterative local</u> particle filter (*Poterjoy 2022*, *QJRMS; Poterjoy 2022, MWR*)

Assimilating obs with non-Gaussian, state-dependent errors.

 Observations: directly measure every 8th variable at Δt = 0.05

•
$$y_i = x_i + \epsilon$$
 for $i = 1, 2, \dots, N_y$



Current approach for specifying $p(\mathbf{y}_t | \mathbf{x}_t^n)$:

Assume $\mathbf{y}_t = H(\mathbf{x}_t^{truth}) + \epsilon_t$, and apply assumptions for distribution of ϵ_t .

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For $\boldsymbol{\epsilon}_t^n = \mathbf{y}_t - H(\mathbf{x}_t^n)$, $p(\mathbf{y}_t | \mathbf{x}_t^n) \approx p(\boldsymbol{\epsilon}_t^n)$, $\approx \mathcal{N}(\boldsymbol{\epsilon}_t^n; \mathbf{0}, \mathbf{R}_t)$. A non-parametric estimate for $p(\mathbf{y}_t | \mathbf{x}_t^n)$:

- Adopt a low-dimensional representation of y_t and x_t from training data using nonlinear manifold learning method (*diffusion maps*; Coifman and Lafon 2006).
- Compute data-driven estimates of p(e_t|x_t) or p(y_t|x_t) using kernel embeddings of conditional distributions (Song et al. 2013; Berry and Harlim 2017).

Results in a matrix representation of $p(\mathbf{y}_t | \mathbf{x}_t)$: To specify likelihood for a given member, find element of matrix that is closest to current \mathbf{y}_t and \mathbf{x}_t^n .

Lorenz example (training time = 40 cycles)

Posterior RMSEs with non-parametric $p(\epsilon_t | \mathbf{x}_t)$

Best Gaussian estimate of $p(\epsilon_t | \mathbf{x}_t)$ (with QC)

The flexibility of data-driven likelihoods opens new research directions.

Another example application:

- We observe the "square" of model variables without knowing this function; i.e., *H* only <u>selects</u> state variables near obs.
- The distribution for ϵ_t is still unknown.
- All <u>5</u> parameters (θ) are unknown (control frequency, amplitude, and coupling between large and small-scale waves).

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$$p(\mathbf{x}_t, \boldsymbol{\theta} | \mathbf{y}_{0:t}) \propto p(\mathbf{x}_t, \boldsymbol{\theta} | \mathbf{y}_{0:t-1}) p(\mathbf{y}_t | \mathbf{x}_t, \boldsymbol{\theta}).$$

Long-term research implications





A new non-Gaussian data assimilation strategy is shown to outperform conventional EnVar used for operational weather prediction.

Early results are encouraging, but the full benefits of non-Gaussian data assimilation still need to be explored.

As a motivating example, we show how likelihoods can be estimated non-parametrically and used for data assimilation with particle filters.







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Kernel embeddings of conditional distributions

We can represents likelihoods using kernel embeddings:

$$p(\mathbf{d}_i|\hat{\mathbf{y}}_j) = \sum_{k=1}^M \mu_{kj} \phi_k(\mathbf{d}_i) q(\mathbf{d}_i)$$

See Song et al. (2009,2013)

$$\mu_{kj} = \sum_{l=1}^{M} \psi_l(\hat{\mathbf{y}}) [\mathbf{C} \tilde{\mathbf{C}}^{-1}]_{kl},$$

$$\mathbf{C}_{lk} = \frac{1}{N} \sum_{j=1}^{N} \phi_l(\mathbf{d}_j) \psi_k(\hat{\mathbf{y}}_j),$$

$$\tilde{\mathbf{C}}_{lk} = \frac{1}{N} \sum_{j=1}^{N} \psi_l(\hat{\mathbf{y}}_j) \psi_k(\hat{\mathbf{y}}_j).$$

where μ_{kj} coefficients determine dependence across **d** and $\hat{\mathbf{y}}$.

For $q(\mathbf{d})$, adopt a kernel estimate:

 Variable bandwidth kernel densities provide non-parametric representation of marginal pdfs.

 $q(\mathbf{d}) = \sum_{k=1}^{N} N(\mathbf{d}_k, \mathbf{B}_k)$, where \mathbf{B}_k is a covariance.

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For basis functions, diffusion maps (Coifman and Lafon 2006) is a reasonable choice:

- Manifold learning method for represent data in lower-dimensional space
- Similar strategy applied by Berry and Harlim (2017)

Constructing basis functions

Example: Data produced from Lorenz (1963) model



Constructing basis functions

Example: Data produced from Lorenz (1963) model



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